## **1** Supplementary Materials

## 2 Supplementary Methods

- 3 Statistical methods
- 4 Survival analysis

5 To accomplish the two study objectives, four variants of the survival Cox model for6 repeated events were fitted.

7 The general formulation of the hazard function for a repeated event survival analysis is8 as follows:

9 
$$\lambda_{i,k}(t;\mathbf{z}_{i,k}(t)) = \lambda_{0k}(t) \alpha_{i} \exp\left(\beta \mathbf{z}_{i,k}(t)\right)$$

10 where  $\lambda_{0k}(t)$  is the ordered event-specific baseline hazard for the  $k^{\text{th}}$  event (k=1, 2 PDs).

11  $\alpha_i$  is the frailty for patient *i* assumed as a Gamma (1,  $\theta$ ) distributed random variable.

12 Parameter  $\theta$  represents patient heterogeneity, by which the higher the  $\theta$ , the greater the

13 heterogeneity. The covariate effect is modeled by the term  $\exp(\beta \mathbf{z}_{ik}(t))$ .  $\mathbf{z}_{ik}(t)$  denotes

14 the covariate vector for the  $i^{th}$  subject and the  $k^{th}$  event and includes treatment with B,

15 potential confounders (possibly time-dependent) and an interaction term for the

16 differential effect of B between first-line and second-line treatment. Since the study

17 design entailed a treatment crossover, B was specified as a time-dependent covariate in

- 18 the statistical analysis.
- 19 Model *a*) derives from assuming the common baseline  $\lambda_0(t)$  for all recurrent events.
- 20 Models a) and b) do not include the interaction term for the differential effect of B.
- 21 Models a), b), and c) fix the frailty term  $\alpha_i$  at 1 for all patients. Not all randomized
- 22 patients experienced PD, and, more importantly, not all patients experiencing the first
- 23 PD were eligible to receive the second-line treatment as planned in the study protocol.

## 24 <u>Calculation of IPCW</u>

The IPCW method aims to reconstruct the complete population, properly weighting the available observations. Therefore, all patient clinical histories post-first PD are, after weighting, representative of the starting population as though the selective withdrawal had never occurred. The weighting formula is as follows:

29 
$$w_{ij} = \prod_{h=0}^{j} \frac{1}{P(A_{ih} = a_{ih} | C_{ih} = c_{ih})}$$

30 where *j* denotes the number of discrete time intervals defined by time-varying covariate for each  $i^{th}$  patient.  $A_{ih}$  indicates the selective withdrawal at the second-line treatment 31 for patient *i* at time point  $t_{ih}$ ,  $C_{ih}$  variables (time-fixed: age at entry, gender, CT 32 regimen, KRAS mutational status, center, study arm, ECOG PS, tumor localization, 33 34 LDH-Lactate dehydrogenase; and time-varying: surgery and toxicity) measured right before each time point  $t_{ih}$  for patient *i*. The equation is a product over all time points 35  $\{t_{ih}\}$  from randomization to discrete time point  $t_{ij}$ , for each patient *i*. Weights are then 36 stabilized using the observed time-fixed covariates and the covariates used for the 37 38 stratified randomization (age at entry, gender, KRAS mutational status, and center). In practice we fit a Cox model for time to second-line treatment ( $A_{ih}$  indicator for 39 40 the selective withdrawal at the second-line treatment) which estimates the probability of selective withdrawal. This model includes all the  $C_{ih}$  variables measured right before 41 each time point  $t_{ih}$  for patient *i*. The IPCW is then the inverse of this probability. To 42 gain in robustness the IPCW is stabilized, *i.e.* calculated as the ratio between the 43 probabilities of selected withdrawals predicted by the Cox model with only the 44 observed time-fixed covariates (age at entry, gender, CT regimen, KRAS mutational 45

- 46 status, center, study arm, ECOG PS, tumor localization, LDH-Lactate dehydrogenase);
- 47 and, in the denominator, the probabilities of selected withdrawals predicted by the Cox
- 48 model with time-fixed and time-varying covariates.