

AMERICAN SOCIOLOGICAL REVIEW

OFFICIAL JOURNAL OF THE AMERICAN SOCIOLOGICAL ASSOCIATION

ONLINE SUPPLEMENT

to article in

AMERICAN SOCIOLOGICAL REVIEW, 2020, VOL. 85

Do Police Brutality Stories Reduce 911 Calls? Reassessing an Important Criminological Finding

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Table S1. Average Daily Calls on New Year's Eve versus other days

| City | New Year's Eve | All Other Days | Number of Years |
|----------------|----------------|----------------|-----------------|
| Baltimore | 1730 | 2034 | 4 |
| Burlington | 16 | 19 | 7 |
| Cincinnati | 1199 | 1378 | 4 |
| Detroit | 852 | 905 | 3 |
| Hartford | 165 | 185 | 3 |
| Las Vegas | 180 | 220 | 3 |
| Los Angeles | 2209 | 2296 | 4 |
| Nashville | 1348 | 1599 | 2 |
| New Orleans | 792 | 872 | 8 |
| Orlando | 660 | 690 | 8 |
| Sacramento | 660 | 747 | 2 |
| San Diego | 1253 | 1413 | 2 |
| Seattle | 305 | 319 | 10 |
| Virginia Beach | 556 | 622 | 2 |

| City | December 24 to 31 | All Other Days | Number of Years |
|----------------|-------------------|----------------|-----------------|
| Baltimore | 1446 | 2039 | 4 |
| Burlington | 14 | 19 | 7 |
| Cincinnati | 1296 | 1588 | 4 |
| Detroit | 782 | 908 | 3 |
| Hartford | 153 | 185 | 3 |
| Las Vegas | 176 | 221 | 3 |
| Los Angeles | 2020 | 2256 | 4 |
| Nashville | 1415 | 1601 | 2 |
| New Orleans | 784 | 873 | 8 |
| Orlando | 616 | 691 | 8 |
| Sacramento | 620 | 749 | 2 |
| San Diego | 1153 | 1417 | 2 |
| Seattle | 274 | 319 | 10 |
| Virginia Beach | 496 | 624 | 2 |

Table S1 shows, for 14 cities with 911 call data readily available online, the average number of daily 911 calls on New Year's Eve compared to all other days (top) and the average number of daily 911 calls for the last seven days of the year compared to all other days. Number of years indicates the number of New Year's Eves in the data. Averages are rounded to the nearest integer. To the extent possible from the provided data fields, calls were subset to deduplicated citizen-initiated police 911 calls by omitting traffic calls, alarm calls, and police-initiated 911 calls, but substantial heterogeneity likely remains between cities in the calls recorded in these data. In every city, calls are lower on the last seven days of the year, so the end-of-year spikes in these Milwaukee data are anomalous.

Table S2. Outlier Changes Signs and Significance of Interaction Terms (Violent Crime)

| Variable | Violent Crime Calls, <i>DPK</i> | Violent Crime Calls, <i>Dropping Final Week</i> |
|---------------------------|---------------------------------|---|
| Weeks Pre-Jude | .019 (.015) | –.002 (.014) |
| Jude Story | –.021 (.065) | –.020 (.066) |
| Weeks Post-Jude | –.177** (.040) | .011 (.041) |
| Weeks Post-Jude (squared) | .003*** (.001) | –.000 (.000) |
| Weeks before Event | 48 | 48 |
| Weeks after Event | 47 | 46 |

* $p < .05$; ** $p < .01$; *** $p < .001$ (two-tailed test).

Table S2 shows estimates of Jude story on violent crime calls from DPK (left) and estimates of same model on same data except for the final week (right) with important differences bolded. The weeks post-Jude coefficients with and without the outlier week are statistically significantly different from one another (Clogg, Petkova, and Haritou 1995):

$$z = \frac{\beta_2 - \beta_1}{\sqrt{(SE\beta_2)^2 + (SE\beta_1)^2}} \approx \frac{0.011 - -0.177}{\sqrt{(0.041)^2 + (0.040)^2}} \approx \frac{0.188}{0.057} \approx 3.3$$

Similarly, the weeks post-Jude (squared) terms are significantly different:

$$z = \frac{\beta_2 - \beta_1}{\sqrt{(SE\beta_2)^2 + (SE\beta_1)^2}} \approx \frac{-0.000 - 0.003}{\sqrt{(0.000)^2 + (0.001)^2}} \approx \frac{0.003}{0.001} \approx 3.0$$

Table S3. Controlling for Outlier Changes Significance and Model Fit

| Variable | Total Calls, <i>DPK</i> | Total Calls, Week 95 Dummy | Total Calls, End-Year Dummies | Total Calls, End-Year Dummies, No Jude Story |
|---------------------------|----------------------------------|-------------------------------|-------------------------------------|---|
| Weeks Pre-Jude | .036*** (.008) | .020* (.008) | .020* (.007) | .013*** (.001) |
| Jude Story | -.009 (.034) | -.008 (.034) | .005 (.034) | |
| Weeks Post-Jude | -.088*** (.021) | .009 (.021) | -.007 (.022) | |
| Weeks Post-Jude (Squared) | .002*** (.000) | -.001 (.000) | -.000 (.000) | |
| Last Week of 2004 | | | .212*** (.030) | .238*** (.029) |
| Last Week of 2005 | | .485*** (.028) | .498*** .028 | .469*** (.029) |
| <i>N</i> | 56,145 | 56,145 | 56,145 | 56,145 |
| BIC | 208325.2 | 208059.6 | 208022.1 | 208007 |

* $p < .05$; ** $p < .01$; *** $p < .001$ (two-tailed test).

Table S3 shows estimates of Jude story on total 911 calls from DPK (left column) and estimates of an otherwise identical model on the same data including dummy parameters for the last weeks of 2004 and 2005 (middle columns) with important differences bolded. Models with both end-of-year dummies have better model fit (with BIC reduced by more than 300). Omitting all parameters associated with the Jude story further improves model fit (right column). The weeks post-Jude coefficients with and without the outlier week are statistically significantly different from one another (Clogg et al. 1995):

$$z = \frac{\beta_2 - \beta_1}{\sqrt{(SE\beta_2)^2 + (SE\beta_1)^2}} \approx \frac{-.0006683 - -.00879}{\sqrt{(.002153)^2 + (.0021469)^2}} \approx \frac{0.008}{0.003} \approx 2.7$$

The weeks post-Jude (squared) terms similarly differ:

$$z = \frac{\beta_2 - \beta_1}{\sqrt{(SE\beta_2)^2 + (SE\beta_1)^2}} \approx \frac{0.0001505 - -.0000294}{\sqrt{(.0000403)^2 + (.0000431)^2}} \approx 3.0$$

Table S4. Outlier Influences Functional Form for Time

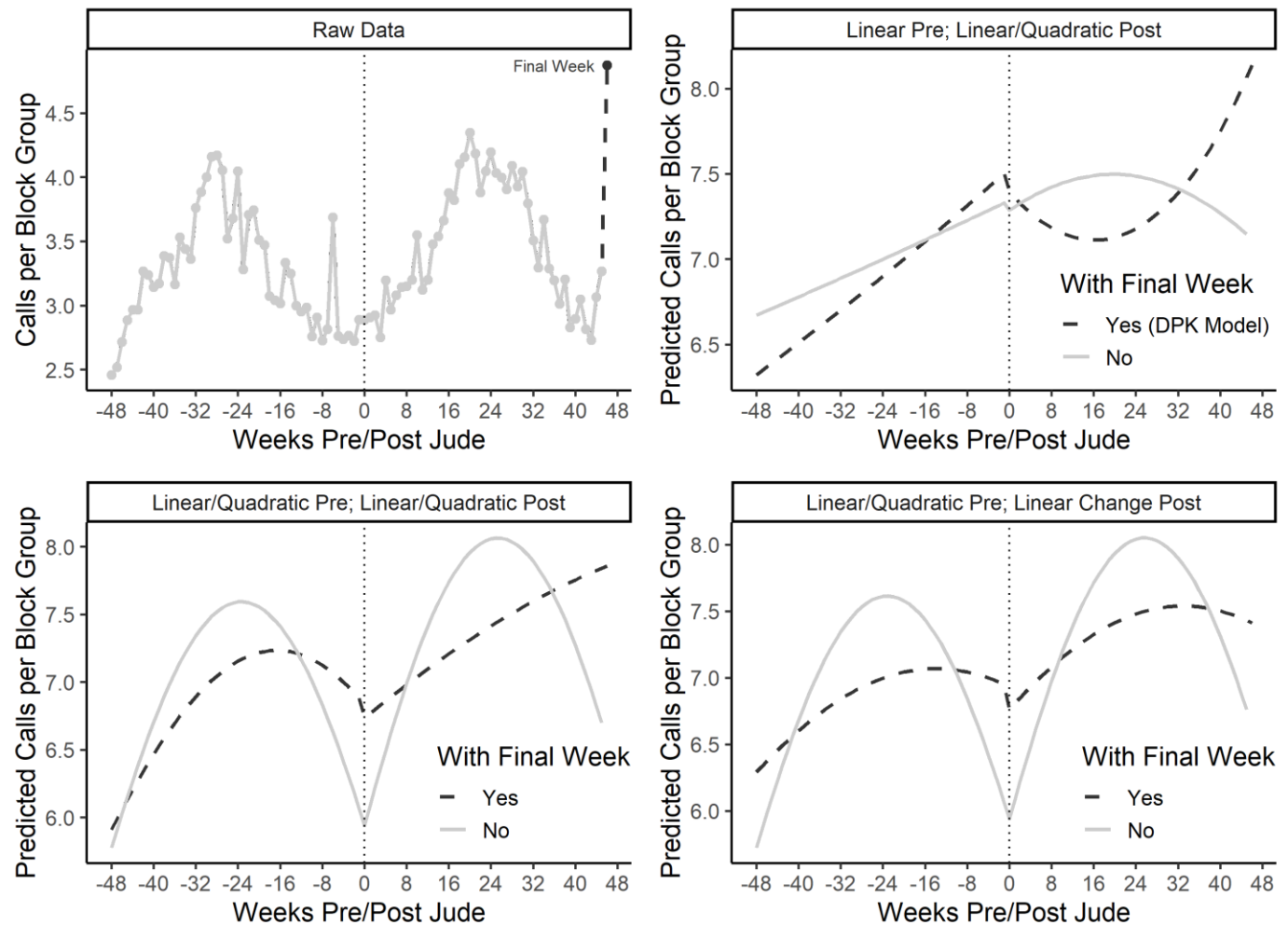
| Model | Drop Week 95 | df | AIC | BIC ($N = \text{CBG} \times \text{Week}$) |
|-------------------------------------|---------------------|-----------|-----------------|---|
| <i>Linear Pre; Linear/Quad post</i> | No | 21 | 208137.6 | 208325.2 |
| Linear/Quad Pre; Linear Post | No | 21 | 208150.2 | 208337.9 |
| Linear/Quad Pre; Linear/Quad Post | No | 22 | 208134.8 | 208331.3 |
| Linear Pre; Linear/Quad Post | Yes | 21 | 205351.1 | 205538.6 |
| Linear/Quad Pre; Linear Change Post | Yes | 21 | 205329 | 205516.4 |
| Linear/Quad Pre; Linear/Quad Post | Yes | 22 | 205330.5 | 205526.9 |

Table S4 reports fit statistics (AIC and BIC) for models with different model specifications of the effect of the Jude story. Smaller AIC and BIC indicate better fit. The italicized specification denotes specification in DPK. Bolded specifications are best fitting for a model estimated on the same data. As in DPK, these specifications always include a change in intercept parameter for the Jude story, although dropping this parameter improves model fit.

Reference

Clogg, Clifford C., Eva Petkova, and Adamantios Haritou. 1995. "Statistical Methods for Comparing Regression Coefficients between Models." *American Journal of Sociology* 100(5):1261–93.

Figure S1. Raw Data and Predicted Values across Specifications



The top left plot of Figure S1 shows the raw call data; the remaining plots show predicted values from different models using Stata's *margins* command, as in DPK. Dashed black lines include the final week; solid gray lines omit the final week. The top right plot shows predicted values from DPK's model's specifications (linear before Jude, linear and quadratic after Jude), and the bottom row uses a symmetric linear/quadratic specification (left) and a linear/quadratic specification with the linear term allowed to change after the story. As in DPK, predicted values are on a different scale than the raw data because Stata's conditional negative binomial fixed-effects model does not estimate the block-group unit intercepts.