

Additional results for moderation analyses that did not yield significant interactions.

### **Anxiety about aging moderation analyses**

**Frequency of contact.** In a multiple regression, gender ( $\beta = -.18, p < .001$ ), years of education ( $\beta = -.16, p < .001$ ), intergenerational contact frequency ( $\beta = -.16, p < .001$ ) and anxiety about aging ( $\beta = .46, p < .001$ ) significantly predicted ageism,  $F(6,403) = 39.63, p < .001$ . This model accounted for 37% of the variance in ageism. The interaction term of intergenerational contact frequency\*anxiety about aging did not significantly predict ageism ( $\beta = -.06, p = .15$ ).

**Quality of contact.** In a multiple regression, gender ( $\beta = -.15, p < .001$ ), years of education ( $\beta = -.16, p < .001$ ), intergenerational contact quality ( $\beta = -.31, p < .001$ ) and anxiety about aging ( $\beta = .39, p < .001$ ) significantly predicted ageism,  $F(6,403) = 51.98, p < .001$ . This model accounted for 44% of the variance in ageism. The interaction term of intergenerational contact quality\*anxiety about aging was not a significant predictor ( $\beta = -.04, p = .39$ ) of ageism.

**Attitudes toward own aging.** In a multiple regression analysis, gender ( $\beta = -.17, p < .001$ ), years of education ( $\beta = -.16, p < .001$ ), and anxiety about aging ( $\beta = .49, p < .001$ ) significantly predicted ageism,  $F(6,404) = 35.88, p < .001$ . This model accounted for 35% of the variance in ageism. However, attitudes toward own aging ( $\beta = -.05, p = .32$ ) and the interaction term of attitudes toward own aging\*anxiety about aging did not significantly predict ageism ( $\beta = .02, p = .68$ ).

### **Age moderation analyses**

**Death anxiety.** In a multiple regression analysis, gender ( $\beta = -.17, p < .001$ ), years of education ( $\beta = -.20, p < .001$ ), and death anxiety ( $\beta = .27, p < .001$ ) significantly predicted ageism,  $F(5,405) = 17.18, p < .001$ . This model accounted for 18% of the variance in ageism. The death anxiety\*age interaction term did not significantly predict ageism ( $\beta = -.01, p = .823$ ).

**Frequency of contact.** In a multiple regression analysis, gender ( $\beta = -.17, p < .001$ ), years of education ( $\beta = -.22, p < .001$ ), and intergenerational contact frequency ( $\beta = -.26, p < .001$ ) significantly

predicted ageism,  $F(5,404) = 18.21, p < .001$ . This model accounted for 18% of the variance in ageism. The interaction term of intergenerational contact frequency\*age did not significantly predict ageism ( $\beta = .02, p = .953$ ).

**Quality of contact.** In a multiple regression, gender ( $\beta = -.13, p < .01$ ), years of education ( $\beta = -.21, p < .001$ ), and intergenerational contact quality ( $\beta = -.45, p < .001$ ) significantly predicted ageism,  $F(5,404) = 38.67, p < .001$ . This model accounted for 32% of the variance in ageism. The interaction term of intergenerational contact quality\*age was not a significant predictor ( $\beta = .05, p = .275$ ) of ageism.

**Attitudes toward own aging.** In a multiple regression analysis, gender ( $\beta = -.15, p < .01$ ), age ( $\beta = -.16, p < .01$ ), years of education ( $\beta = -.18, p < .001$ ), and attitudes toward own aging ( $\beta = -.26, p < .001$ ) significantly predicted ageism,  $F(5,405) = 18.80, p < .001$ . This model accounted for 19% of the variance in ageism. However, the interaction term of attitudes toward own aging\*age did not significantly predict ageism ( $\beta = .08, p = .10$ ).